Ensemble Kalman filter and related filters

Schedule

- ▶ 16 Jan: Gaussian processes (Jo Eidsvik)
- 23 Jan: Hands-on project on Gaussian processes (Team effort, work in groups)
- 30 Jan: Latent Gaussian models and INLA (Jo Eidsvik)
- 6 Feb: Hands-on project on INLA (Team effort, work in groups)
- 12-13 Feb: Template model builder. (Guest lecturer Hans J Skaug)
- 20 Feb: Ensemble Kalman filter and related filters (Jo Eidsvik)
- 27 Feb: Ensemble Kalman filter and related filters (Henning Omre)

- 6 Mars: Discrete random processes (Håkon Tjelmeland)
- 13 Mars: No Lecture!
- 20 Mars: Discrete random processes (Håkon Tjelmeland)
- 10 April: Dim red?..

Bayes formula

$$p(\mathbf{x}|\mathbf{y}) = \frac{p(\mathbf{x}, \mathbf{y})}{p(\mathbf{y})} = \frac{p(\mathbf{x})p(\mathbf{y}|\mathbf{x})}{\int p(\mathbf{x})p(\mathbf{y}|\mathbf{x})d\mathbf{x}}$$

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 $p(\mathbf{x})$ is prior. $p(\mathbf{y}|\mathbf{x})$ is likelihood.

Example of such a model

$$m{x} \sim N(m{\mu}_x, m{\Sigma}_x)$$

$$oldsymbol{y}|oldsymbol{x}\sim \mathcal{N}(oldsymbol{H}oldsymbol{x},oldsymbol{T}), \quad oldsymbol{y}=oldsymbol{H}oldsymbol{x}+\epsilon, \ \epsilon\sim \mathcal{N}(oldsymbol{0},oldsymbol{T})$$

 $p(\mathbf{x}|\mathbf{y})$ is Gaussian with

$$E(\mathbf{x}|\mathbf{y}) = \mu_x + \boldsymbol{\Sigma}_x \boldsymbol{H}' [\boldsymbol{H} \boldsymbol{\Sigma}_x \boldsymbol{H}' + \boldsymbol{T}]^{-1} (\mathbf{y} - \boldsymbol{H} \mu_x),$$

Var $(\mathbf{x}|\mathbf{y}) = \boldsymbol{\Sigma}_x - \boldsymbol{\Sigma}_x \boldsymbol{H}' [\boldsymbol{H} \boldsymbol{\Sigma}_x \boldsymbol{H}' + \boldsymbol{T}]^{-1} \boldsymbol{H} \boldsymbol{\Sigma}_x.$

Mean is linear in conditioning variable.

Variance is not dependent on conditioning variable, only correlations and variances.

Sequential Bayesian assimilation



 $p(\mathbf{x}_1), p(\mathbf{x}_t | \mathbf{x}_{t-1}), p(\mathbf{y}_t | \mathbf{x}_t), j = 2, 3, \dots, T.$

Dynamic model

Process model is described by:

$$p(\boldsymbol{x}_t|\boldsymbol{x}_{t-1},\ldots,\boldsymbol{x}_1)=p(\boldsymbol{x}_t|\boldsymbol{x}_{t-1}),$$

This could be a differential equation, or it could be a simple linear process, or even a static process ($\mathbf{x}_t = \mathbf{x}_{t-1}$).

The data gathering process is described via the likelihood:

$$p(\mathbf{y}_t|\mathbf{x}_t,\ldots,\mathbf{x}_1,\mathbf{y}_{t-1},\ldots,\mathbf{y}_1) = p(\mathbf{y}_t|\mathbf{x}_t)$$

This could also be nonlinear, or it could represent picking a subset of variables (with noise).

General formula

Filtering, solution:

$$p(\mathbf{x}_t|\mathbf{y}_1,\ldots,\mathbf{y}_{t-1}) \propto \int p(\mathbf{x}_t|\mathbf{x}_{t-1})p(\mathbf{x}_{t-1}|\mathbf{y}_1,\ldots,\mathbf{y}_{t-1})d\mathbf{x}_{t-1}.$$

$$p(\mathbf{x}_t|\mathbf{y}_1,\ldots,\mathbf{y}_t) = \frac{p(\mathbf{x}_t|\mathbf{y}_1,\ldots,\mathbf{y}_{t-1})p(\mathbf{y}_t|\mathbf{x}_t)}{p(\mathbf{y}_t|\mathbf{y}_1,\ldots,\mathbf{y}_{t-1})}$$

Note Markov assumption in process, and conditionally independent data.

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Kalman filter

For a Gaussian prior $\mathbf{x}_1 \sim N(\boldsymbol{\mu}_1, \boldsymbol{\Sigma}_1)$, linear Gaussian dynamic model $\mathbf{x}_t | \mathbf{x}_{t-1} \sim N(\mathbf{G}_t \mathbf{x}_{t-1}, \mathbf{Q})$, and linear Gaussian likelihood $\mathbf{y}_t | \mathbf{x}_t \sim N(\mathbf{H}_t \mathbf{x}_t, \mathbf{R})$, there exists an exact recursion for the filtering distribution: $p(\mathbf{x}_t | \mathbf{y}_1, \dots, \mathbf{y}_t) = N(\mathbf{m}_t, \mathbf{V}_t)$.

Initialization:

$$\boldsymbol{\mu}_1 = \boldsymbol{\mu}_1, \quad \boldsymbol{\Sigma}_1 = \boldsymbol{\Sigma}_1,$$

• Recursive updating for $j = 1, \ldots, T$:

$$S_t = H_t \Sigma_t H_t^t + R,$$

$$K_t = \Sigma_t H_t^t S_t^{-1},$$

$$m_t = \mu_t + K_t (y_t - H_t \mu_t),$$

$$V_t = \Sigma_t - K_t H_t \Sigma_t.$$

$$\Sigma_{j+1} = G_t V_t G_t^t + Q$$

$$\mu_{j+1} = G_t m_t$$

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Non-Gaussian or non-linear

In other situations there is usually no exact solution to the filtering distribution. Approximations:

- Extended Kalman filter (EKF) : linearization
- Unscented Kalman filter (UKF) : design points and 'numerical' integration
- Ensemble Kalman filter (EnKF) : Monte Carlo samples and linear updates

Particle filter (PF) : Simulation and likelihood weighting.

Algorithms

Summary of some filtering methods ; pros and cons.

Criterion	EKF	UKF	EnKF	PF
Analytic conditioning	V	V	V	
MC based			V	V
Non-linear	w	V	V	V
Scales with dim.	V		V	
Reliable UQ		W		w

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Ensemble Kalman filter (EnKF)

- Monte Carlo based data assimilation
- Assimilation based on linear update

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(Evensen, 1994, Evensen, 2009).

EnKF

- ▶ Initial: Independent prior samples $\mathbf{x}_0^{b,a} \sim p(\mathbf{x}_0), b = 1, \dots, B$.
- ► Iterate for samples b = 1,..., B and time steps t = 1,..., T: Forecast variables (could be non-linear, black-box solver):

$$\boldsymbol{x}_{t}^{b,f} = \boldsymbol{g}(\boldsymbol{x}_{t-1}^{b,a}; \boldsymbol{\epsilon}_{t}^{b}),$$

Forecast data (could be non-linear, black-box solver):

$$\boldsymbol{y}_t^b = \boldsymbol{h}(\boldsymbol{x}_t^{b,f}; \boldsymbol{\delta}_t^b),$$

Assimilate:

$$\begin{aligned} \mathbf{x}_t^{b,a} &= \mathbf{x}_t^{b,f} + \hat{\mathbf{K}}_t(\mathbf{y}_t - \mathbf{y}_t^b).\\ \hat{\mathbf{K}}_t &= \hat{\mathbf{\Sigma}}_{xy,t} \hat{\mathbf{\Sigma}}_{y,t}^{-1}. \end{aligned}$$

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EnKF update

$$\boldsymbol{x}_t^{b,a} = \boldsymbol{x}_t^{b,f} + \hat{\boldsymbol{\Sigma}}_{xy,t}\hat{\boldsymbol{\Sigma}}_{y,t}^{-1}(\boldsymbol{y}_t - \boldsymbol{y}_t^b).$$

This is a regression problem.

$$\begin{split} \hat{\boldsymbol{\Sigma}}_{y,t} &= \frac{1}{B} \sum_{b=1}^{B} (\boldsymbol{y}_{t}^{b} - \bar{\boldsymbol{y}}_{t}) (\boldsymbol{y}_{t}^{b} - \bar{\boldsymbol{y}}_{t})', \quad \bar{\boldsymbol{y}}_{t} = \frac{1}{B} \sum_{b=1}^{B} \boldsymbol{y}_{t}^{b} \\ \hat{\boldsymbol{\Sigma}}_{xy,t} &= \frac{1}{B} \sum_{b=1}^{B} (\boldsymbol{x}_{t}^{b,f} - \bar{\boldsymbol{x}}_{t}) (\boldsymbol{y}_{t}^{b} - \bar{\boldsymbol{y}}_{t})', \quad \bar{\boldsymbol{x}}_{t} = \frac{1}{B} \sum_{b=1}^{B} \boldsymbol{x}_{t}^{b,f}, \end{split}$$

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Univariate example - samples

$$x^{b} \sim p(x), \quad y^{b} = x^{b} + N(0, 5^{2})$$



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Univariate example - regression fit



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Univariate example - observation



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Univariate example - analysis or update step



Univariate example - prior and posterior



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Univariate - one-step problem

Implement the Ensemble Kalman filter (one time step only) on the following model, and compare with the exact solution for p(x|y). Plot solutions.

$$x \sim N(0,1), \quad y = x^{2q+1} + N(0,0.05^2)$$

Data is y = 2.

Set q = 0. Study performance of the EnKF for different ensemble sizes B = 20, 100.

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• Study performance of the EnKF for different q = 1, 2, ...

EnKF, Gauss-linear likelihood

- ▶ Initial: Independent samples $\mathbf{x}_0^{b,a} \sim p(\mathbf{x}_0), b = 1, \dots, B$.
- ► Iterate for samples b = 1,..., B and time steps t = 1,..., T: Forecast variables (could be non-linear, black-box solver):

$$\boldsymbol{x}_{t}^{b,f} = \boldsymbol{g}(\boldsymbol{x}_{t-1}^{b,a}; \boldsymbol{\epsilon}_{t}^{b}),$$

Linear likelihood:

$$\boldsymbol{y}_t^b = \boldsymbol{H}\boldsymbol{x}_t^{b,f} + N(0,\boldsymbol{T})$$

Assimilate :

$$\boldsymbol{x}_{t}^{b,a} = \boldsymbol{x}_{t}^{b,f} + \hat{\boldsymbol{\Sigma}}_{t}\boldsymbol{H}^{'}(\boldsymbol{H}\hat{\boldsymbol{\Sigma}}_{t}\boldsymbol{H}^{'} + \boldsymbol{T})^{-1}(\boldsymbol{y}_{t} - \boldsymbol{y}_{t}^{b}).$$

Estimation of Σ_t

Standard approach:

$$\hat{\boldsymbol{\Sigma}}_t = rac{1}{B}\sum_{b=1}^B (\boldsymbol{x}_t^{b,f} - ar{\boldsymbol{x}}_t) (\boldsymbol{x}_t^{b,f} - ar{\boldsymbol{x}}_t)', \quad ar{\boldsymbol{x}}_t = rac{1}{B}\sum_{b=1}^B \boldsymbol{x}_t^{b,f}$$

Gives less Monte Carlo error than straightforward estimator for Kalman gain.

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Gaussian random field



Use B = 100 ensembles (realizations) to estimate the covariance matrix.

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Estimation of covariance matrix



There is lots of Monte Carlo error in the estimated covariance matrix and Kalman gain. Numerous tricks try to resolve this: inflation, localization, etc. Also, since the Kalman gain is estimated from data, there is coupling over many time steps which gives challenges. Resampling from a Gaussian approximation might solve this. (More details next week.).

Spatial AR(1) model

 $\mathbf{x}_0 \sim N(0, \mathbf{\Sigma}),$

$$\mathbf{x}_{t} = \rho \mathbf{x}_{t-1} + N(0, (1 - \rho^{2})\mathbf{\Sigma}), \ t = 1, \dots, T$$

$$y_t = x_t + N(0, \tau^2 I), t = 1, ..., T$$

 $T = 10, 50 \times 50$ grid.

EnKF is run with B = 5000 ensemble members.

Ensemble Kalman filter and related filters

Spatial AR(1): KF pred





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Spatial AR(1): EnKF pred





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Spatial AR(1): MSE and coverage



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KF (o), EnKF (+).

Spatial AR(1): EnKF pred (B = 500)



(Average coverage for EnKF is 0.38 for nominal 0.90.)

Univariate - AR(1) time series

Implement the Ensemble Kalman filter on the following model, and compare with the exact Kalman filter solution. Plot solutions (mean with 90 % prediction intervals). Study performance for B = 20,100,1000 in the EnKF.

$$x_1 \sim N(0,1),$$

$$x_t = 0.9x_{t-1} + N(0, 1 - 0.9^2), \ t = 2, \dots, 50$$

$$y_t = x_t + N(0, 0.2^2), t = 1, \dots, 50$$

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Main points

- ► Filtering methods, sequential data assimilation.
- EnKF is ensemble based, with linear update.
- EnKF has been successful in many applications, but often shows underestimation of variability.

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• EnKF is complicated in models with very nonlinear likelihoods.